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Journal article

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The authors declare that they have no relationship with companies or manufacturers who will benefit from the results of the present study. The results of the present study do not constitute endorsement by ACSM. The results of the study are presented clearly, honestly, and without fabrication, falsification, or inappropriate data manipulation.

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ABSTRACT

Introduction Long-term effects of physical activity and TV viewing on mortality have been inferred from observational studies. The associations observed do not allow inferences about the effects of population interventions and could be subject to bias due to time-varying confounding.

Methods Using data from the Australian Diabetes, Obesity and Lifestyle Study, collected in 1999-2000 (T0), 2004-05 (T1), and 2011-12 (T2), we applied the parametric g-formula to estimate cumulative risks of death under hypothetical interventions on physical activity and/or TV viewing determined from self-report, while adjusting for time-varying confounding.

Results In the 6,377 participants followed for 13 years from 2004-05 to death or censoring in 2017, 781 participants died. The observed cumulative risk of death was 12.2%. The most effective hypothetical intervention was to increase weekly physical activity to >300 minutes (RR=0.66, 0.46 to 0.86 compared with a 'worst-case' scenario; and RR=0.83, 0.73 to 0.94 compared with no intervention). Reducing daily TV viewing to <2 hours in addition to physical activity interventions did not show added survival benefits. Reducing TV viewing alone was least effective in reducing mortality (RR=0.85, 0.60 to 1.10 compared with the worst-case scenario; and RR=1.06, 0.93 to 1.20 compared with no intervention).

Conclusion Our findings suggested that sustained interventions to increase physical activity could lower all-cause mortality over a 13-year period and there might be limited gain from intervening to reduce TV viewing time in a relatively healthy population.

Keywords: time-varying confounding, hypothetical interventions, g-formula, cohort study

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Introduction

Both insufficient physical activity (i.e., not meeting physical activity recommendations) and sedentary behaviour (time spent sitting, as distinct from lack of physical activity) contribute to risk of chronic disease and mortality. In the absence of evidence from randomized trials to quantify the long-term effects of changes in physical activity and sedentary behaviour, understanding how they are jointly related to mortality could be enhanced by better exploiting data from observational studies (1).

Insufficient physical activity and time spent in sedentary behaviours, particularly television (TV) viewing, have been associated with higher all-cause mortality in observational studies (2, 3). These studies have typically measured exposures and confounders at a single time point, so did not assess the possible impact of exposure changes over follow-up. We have previously highlighted (4) that in studies that used data from multiple time points, conventional regression analyses can be problematic in the presence of time-varying confounding when the values of confounding variables are influenced by past exposures (e.g. sedentary behaviour affects adiposity, which in turn affects physical activity at the next time point) (5, 6). When there is time-varying confounding, conditioning on confounders (e.g. adiposity) that also lie in a causal pathway in standard regression models can produce biased estimates (see Figure, Supplemental Digital Content 1, which illustrates an example of time-varying confounding affected by prior exposure, <http://links.lww.com/MSS/C89>) (7). Alternative methods such as inverse probability weighting of marginal structural models have been used to estimate causal effects of physical activity while adjusting for such time-varying confounding (8-12). No published studies on

sedentary behaviour with multiple observation points have accounted for time-varying confounding.

Insufficient physical activity and sedentary behaviour can be viewed as separate risk factors with distinct sociodemographic and behavioural contexts and correlates (13). Our aim was to estimate the effects of single or joint hypothetical interventions for insufficient physical activity and a common leisure-time sedentary behaviour, TV viewing, on all-cause mortality over an approximate 13-year period, while accounting for time-varying confounding, using the parametric g-formula. We used the parametric g-formula because it allows estimation of the causal effects of complex population interventions, which could inform policy more directly compared with a typical exposure effect (14).

Methods

Study population

The Australian Diabetes, Obesity and Lifestyle Study (AusDiab) is a population-based cohort study conducted in the six states and the Northern Territory of Australia. Details about the cohort have been described (15). Briefly, participants aged at least 25 years were recruited in 1999-2000 (T0), then followed up in 2004-05 (T1), and 2011-12 (T2). Each data collection involved an initial household interview, followed by a biomedical examination and the administration of questionnaires. In the present study, we used T1 (2004-05) as baseline in order to have information on pre-baseline exposure and confounder history. Participants who attended T1 data collection (n=6,400) were included in this analysis. Participants who were pregnant (n=23) at data collections were excluded, which left 6,377 participants eligible for the analysis. The study

was approved by the Ethics Committee of the International Diabetes Institute and all participants provided informed consent.

Exposure measurements

Self-reported frequency and duration of leisure-time physical activity during the previous week was measured using the Active Australia Survey (16). The questions have been shown to have good reliability and validity (16). Physical activity consisted of walking for recreation or transport, moderate-intensity and vigorous-intensity physical activity at leisure-time. Total weekly recreational physical activity time was calculated as the sum of the time spent walking continuously for at least 10 minutes, time performing moderate physical activity, and double the time spent in vigorous physical activity ($Time_{LTPA} = Time_{walk} + Time_{moderate} + 2 \times Time_{vigorous}$) (16). The total time of weekly physical activity during leisure-time was later used to simulate hypothetical interventions.

Participants were asked to self-report time spent watching TV or videos in the last 7 days, but to exclude the time that this occurred while performing other activities such as preparing a meal or doing other household chores. This method has been shown to provide reliable and valid estimates of TV viewing time among adults (17). Average daily TV viewing hours was calculated.

Confounder measurements

Information on demographic attributes (sex, baseline age, country of birth, and level of education), self-reported history of health conditions (high cholesterol, high blood pressure, heart

disease, and diabetes), and smoking status was obtained by an interviewer-administered questionnaire (15). Total cholesterol, hypertension and diabetes status were also measured in the biomedical examinations. However, we used the self-reported history of diagnosis, assuming that awareness of prior health conditions would have more influence on an individual's lifestyle behaviours. Quintiles of an area-based index of relative socioeconomic advantage and disadvantage was calculated based on participants postcode of residence (18). Alcohol and dietary intake were assessed using a self-administered, validated food frequency questionnaire. Mediterranean diet score was computed and used as a measure of overall diet quality (19). General health was assessed using the self-administered SF-36 questionnaire. Waist circumference was measured by trained staff (15). We used waist circumference instead of BMI because there is evidence it is a stronger predictor of all-cause mortality than BMI (20).

Death ascertainment

Vital status and date of death were determined by linkage to the Australian National Death Index. Participants were followed until the date of death or administrative end of follow-up on 17 April 2017.

Hypothetical interventions

We considered the following hypothetical interventions at T1 and T2, based on guidelines for physical activity (21) and the associations between TV viewing time and metabolic biomarkers (22): increasing weekly physical activity to sufficient (i.e. 150 to 300 minutes) if insufficiently active (i.e. <150 minute); increasing weekly physical activity to optimal (i.e. >300 minutes) for all participants; reducing daily TV viewing to <2 hours for all participants; increasing weekly

physical activity to sufficient if insufficiently active and reducing daily TV viewing to <2 hours; and, increasing weekly physical activity to >300 minutes and reducing daily TV viewing to <2 hours for all participants.

In addition, for comparison, we considered a no-intervention scenario in which physical activity level and TV viewing time were allowed to evolve naturally (typically referred to as the ‘natural course’), and a scenario where weekly physical activity decreased to less than 30 minutes and daily TV viewing increased to 4 hours or more for all participants (i.e. worst-case scenario).

Statistical analysis

We used the parametric g-formula to estimate the 13-year cumulative risk of death under various hypothetical interventions on physical activity and/or TV viewing. The parametric g-formula is a generalization of standardization for time-varying exposures and confounders and can be used to estimate the standardized risk of death for hypothetical interventions under the assumptions of no unmeasured confounding, no measurement error and no model misspecification (6). The standardized risk is estimated by a weighted average of the risks of death conditional on the given intervention and the observed confounder history. The weights are probability distribution functions of the time-varying confounders estimated using parametric regression models. The weighted average is approximated through Monte Carlo simulation (23). We implemented the parametric g-formula in two steps. First, parametric models were fitted to model conditional probabilities of physical activity, TV viewing, and each of the following time-varying confounders in the order listed: self-reported history of high cholesterol (yes; no), high blood pressure (yes; no), heart disease (yes; no), and diabetes (yes; no), self-reported general health

status (excellent; very good; good; fair; poor), waist circumference (normal: <94cm [male] or <80cm [female]; increased risk: 94cm to <102cm [male] or 80cm to <88 cm [female]; greatly increased risk: ≥102cm [male] or ≥88 [women] (24)), Mediterranean diet score (0–3; 4–6; 7–9 (19)), smoking status (never; former; current), and alcohol intake (gram/day: 0 [Male & Female]; 1–39 [Male]/1–19[Female]; 40–59[Male]/20–39[Female]; 60+[Male]/40+[Female] (25)).

The models also included the following time-fixed confounders: sex (male; female), baseline age (years), quintiles of an area-based index of relative socio-economic advantage and disadvantage (18), country of birth (Australia or New Zealand; others), and level of education (university or technical institution; completed high school; some high school; primary/never attended school). See Table, Supplemental Digital Content 2, for details of models, <http://links.lww.com/MSS/C90>.

These models were then used to simulate risk of death while setting physical activity and TV viewing to a specified intervention level in a Monte Carlo sample of the same size: 1) T0 and T1 confounder values were retained for all participants; T1 physical activity and TV viewing values were set to a specific level if part of an intervention; 2) risk of death before T2 was simulated; 3) for participants simulated to remain alive at T2: physical activity and TV viewing were set to a specific level if part of an intervention, T2 values of confounders were simulated by comparing the predicted probability of the confounder value to a value randomly drawn from a standard uniform distribution, and risk of death from T2 to the end of follow-up was simulated; 4) cumulative risk of death (i.e. 13-year risk) was calculated as:

$$P_{13\text{-year}} = P_{\text{death before T2}} + (1 - P_{\text{death before T2}}) P_{\text{death after T2}}$$

For each hypothetical intervention, we compared the estimated 13-year risk of death with the risk under the natural course (i.e. no-intervention scenario) and the risk under the worst-case scenario by calculating the risk ratios (RR) and risk differences (RD). We conducted the analyses separately in female and male participants to examine the possibility of effect heterogeneity by sex. We also compared simulated risk of death under the natural course with the observed risk as an informal validation of correct gross model specification.

Multiple imputation by chained equations (MICE) was used to impute missing data (due to missing response to the questionnaire, or missing T2 attendance for those who were still alive at T2) under the assumption that data were missing at random, i.e. the probability of data being missing did not depend on the unobserved data, conditioning on the observed data (26). For each hypothetical intervention, point estimates were averaged over 40 imputed datasets; For the main analysis, 500 bootstrap samples were drawn for each imputed data set to estimate the standard errors and 95% confidence intervals were calculated using Rubin's rule (27, 28); for sensitivity analyses, 200 bootstrap samples were used.

For comparison with a conventional approach, Cox regression with age as the time scale was used to estimate hazard ratios for mortality associated with baseline TV viewing and physical activity, adjusting for baseline confounders.

Statistical analyses were performed using Stata version 14.2 (StataCorp, Texas, USA), and Stata version 15 on the University of Melbourne's high performance computing platform (Spartan).

Results

A total of 6,377 participants (54.7% female) were eligible. During 13 years (73,518 person years) of follow-up, 781 participants died (373 pre-T2 and 408 post-T2). Of participants who were alive at T2 (n=6,004), 20% did not attend T2 data collection. Participants who attended T2 were overall younger, from more socio-economically advantaged areas, and had higher educational qualifications than participants who were alive but did not attend T2; a higher proportion of them reported good to excellent health in general, and no history of health conditions at T1 (see Table, Supplemental Digital Content 3, characteristics of participants who were alive but did not attend T2, <http://links.lww.com/MSS/C91>).

Table 1 shows the characteristics of eligible participants at baseline (T1), and the potential time-varying confounders pre- and post-baseline. Mean age at baseline was 56.5 years. Three quarters (75.9%) were born in Australia or New Zealand and 40.2% had tertiary education. At baseline, more than half of the sample were sufficiently active (57%) or watched less than 2 hours of TV (54%) (Table 1). Active participants tended to spend less time watching TV daily, although the differences were not large (Figure 1).

Table 2 shows the 13-year risks of death under various hypothetical interventions. The simulated 13-year risk of death under no intervention (12.1%) was very similar to the observed risk (12.2%), indicating that the models were correctly specified overall. The hypothetical intervention that reduced 13-year risk of death the most was to improve physical activity to >300 mins/week (RR=0.83, 0.73 to 0.94 compared with the natural course; and RR=0.66, 0.46 to 0.86 compared with the worst-case scenario), followed by improving physical activity to 150-300

mins/week for insufficiently active participants (RR=0.92, 0.82 to 1.01 compared with the natural course; and RR=0.73, 0.52 to 0.94 compared with the worst-case scenario). The average percentages of participants who needed to improve their physical activity were 65.2% and 42.1%, respectively for the two interventions. The intensive physical activity intervention would have prevented 20 deaths (CI: 7 to 33 deaths) per 1000 people in a 13-year period compared with not intervening. Reducing daily TV viewing to < 2 hours alone was the least effective intervention for lowering mortality (RR=1.06, 0.93 to 1.20 compared with the natural course; and RR=0.85, 0.60 to 1.10 compared with the worst-case scenario). Reducing daily TV hours jointly with any of the physical activity interventions required more people changing their behaviours (average of 80.7% and 68.2%, respectively), while not lowering the risk further.

Table 3 shows the 13-year risk of death in male and female participants under the natural course, the worst-case scenario, and the joint intensive intervention. The effect of hypothetical interventions on mortality (i.e. risk ratios) appeared to be similar for male and female participants. However, population risk difference was larger in males than in females because of higher absolute risks under the natural course.

We assumed correct ordering of exposures and time-varying confounders in our models. Our sensitivity analysis showed that results were robust to various modelling orders (see Table, Supplemental Digital Content 4, estimated risk of death under various modelling orders of time-varying covariates, <http://links.lww.com/MSS/C92>).

We found that the usual method of analysis, which used only baseline data in a Cox regression model underestimated the benefit of sustained higher physical activity compared with the g-formula, but the effect of TV viewing on all-cause mortality estimated from the g-formula was similar to the effect estimated from the Cox regression (Figure 2).

Discussion

Our results suggest that in this cohort of adults, mortality could have been lowered by sustained interventions that increased physical activity. The intervention that appeared most effective to reduce mortality compared with no intervention was to increase weekly physical activity to >300 minutes (the intensive physical activity intervention), followed by increasing physical activity to 150-300 minutes/week in people who were insufficiently active (the moderate physical activity intervention). Interventions that reduced TV viewing time alone or in addition to physical activity interventions did not show added mortality benefits.

Although the intensive physical activity intervention was the most effective in reducing mortality, it required more participants to modify their behaviour to achieve the change (on average, 65% of participants needed to modify their physical activity levels at each time point), compared with the moderate physical activity intervention (42% on average needs to change). A systematic review found that relative reduction in all-cause mortality associated with higher physical activity was greater for females than for males (29), the effects of the hypothetical interventions on relative reduction in mortality were similar for females and males in our study. It should be noted that in real life incomplete adherence is likely, and our estimates correspond to the best-case scenario.

Like other analyses of observational data, these estimates are based on the assumptions of no unmeasured confounding, no measurement error, and no model misspecification. We cannot exclude the possibility of unmeasured confounding despite adjusting for several important confounders. Self-reported time spent in physical activity and TV viewing are subject to measurement error. However, the questionnaires used in our study were previously shown to have good reliability and acceptable validity for estimates of the true exposure level (16, 17). We acknowledge the possibility of misspecification of parametric models and functional forms of the past covariate history included as independent variables. However, we were able to closely reproduce the observed risk of death under the natural course, which is a necessary condition for no overall model misspecification under no intervention. The parametric g-formula requires fitting multiple models, therefore it may be more sensitive to violations of the above assumptions, as violation in one model may accumulate through the others (23). The parametric g-formula is subject to the ‘g-null paradox’, i.e. the null hypothesis, (in our case, this is that interventions on physical activity and TV viewing have no effect on all-cause mortality), even if true, will be rejected in a large enough sample because the estimated value of the g-formula for the outcome generally depends on the exposure history (30). However, in practice, the g-null paradox is of less concern compared with typical random variability (31). Finally, the imputation of missing data relied on the missing-at-random assumption, which was reasonably plausible given the extensive risk factor information collected at each time point.

Current public health guidelines recommend minimizing sedentary behaviour and doing at least 150 mins/week of moderate-to-vigorous-intensity physical activity, or 300 mins/week for

additional health benefit (32-34). These recommendations are mainly based on studies of associations between exposures at a single time point and risk of health outcomes such as cardiovascular health and cancer (33). Our study, on the other hand, estimated the potential impact on mortality had these two risk factors been altered by sustained population interventions. This is the key strength of our study, because it is rarely feasible to estimate such causal effects for a generally healthy population through randomized controlled trials (1). Our finding demonstrated that using a single measurement of physical activity is likely to underestimate the protective effects of physical activity. This may stimulate additional public health expenditure into physical activity promotion. Health promotion programmes frequently incorporate physical activity promotion into programmes to address obesity prevention or reduction. Our research (which accounts for obesity-related time-varying confounding) highlights that physical activity itself is important for longevity. Although other g-methods such as inverse probability weighted marginal structural models could also overcome the bias from time-varying confounding affected by past exposure through generating a pseudo-population in which exposure is independent of confounders, the parametric g-formula has the advantage of generating counterfactual outcomes under different exposure scenarios that involve multiple interventions such as increasing physical activity *and* reducing TV time (7).

Previous findings from the AusDiab study reported that watching ≥ 4 hours of TV daily was associated with higher all-cause mortality (35). Our Cox model showed a weaker association in the same direction between TV viewing time at T1 and all-cause mortality (Figure 2). This could be partly because the previous study used T0 as baseline, whereas we used T1 as baseline. Our sample was smaller due to loss to follow-up between T0 and T1, and healthier. The prevalence of

self-reported excellent and very good health status was slightly lower in participants than in the general Australian population (36). However, the participants were more physically active than the general Australian population of the same age – about 50% at T0 reported sufficient levels of activity (see Table, Supplemental Digital Content 5, pre-baseline characteristics of participants by baseline attendance status, <http://links.lww.com/MSS/C93>) compared with about 40% for the population (37). Thus, our findings might be most applicable to physically active people. In our sample where daily TV viewing hours were already below two hours for more than half of the participants, we estimated no further survival benefit by intervening on this exposure. Over the 12 years between T0 (1999/2000) and T2 (2011/2012), there was an expansion of television viewing options, and other domestic entertainment and screen-based technologies, which may have reduced the relevance of our exposure variable. Although our estimates are not directly comparable to results from studies using conventional regression approaches, our findings and those of studies using regression approaches suggest protective effects of physical activity on mortality (29). Furthermore, we found that using only baseline data could underestimate the potential benefit of long-term physical activity.

Although we used repeatedly measured exposure data, the analyses would have benefited from more time points at regular intervals, which are more representative of sustained interventions over time. We coarsened the time spent in physical activity and TV viewing into categories relevant to current public health guidelines. This may affect the interpretation of our findings because of multiple versions of treatment (38). For example, our hypothetical intervention, “increasing physical activity to > 300 mins/week” can be achieved by increasing physical activity to 301 minutes or to 400 minutes through increasing activity duration or intensity over a

week. Our estimates can be interpreted as a weighted average of the effects of the different versions, weighted by the probability of each version naturally arising within the population (38, 39). It should be noted that our estimates may not be generalizable to populations with different distributions of physical activity and TV viewing level. Results from the Australian National Health Surveys showed that the percentage of Australian adults with sufficient physical activity (i.e. ≥ 150 mins/week) remained low from 1989 to 2011 (39% in 1989 to 41% in 2011) (37). The hypothetical interventions we considered may have a greater benefit on lowering mortality in the general population than in our sample where close to 60% can be classified as ‘sufficiently active’.

In conclusion, our findings suggest that sustained interventions on physical activity could lower all-cause mortality over a 13-year period, and that there might be limited gain from intervening on TV viewing time in a relatively healthy population.

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YY, AMH, PAD, BML, and DRE designed the study. YY performed the statistical analysis with support from EJW. YY, AMH, PAD, BML, and DRE drafted the manuscript. PAG, ELMB, NO, and DWD contributed to the data interpretation and provided critical feedback for each draft. All authors read and approved the final manuscript.

Conflict of interest

The authors declare that they have no relationship with companies or manufacturers who will benefit from the results of the present study. The results of the present study do not constitute endorsement by ACSM. The results of the study are presented clearly, honestly, and without fabrication, falsification, or inappropriate data manipulation.

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References

1. Lynch BM, Leitzmann MF. An Evaluation of the Evidence Relating to Physical Inactivity, Sedentary Behavior, and Cancer Incidence and Mortality. *Current Epidemiology Reports*. 2017;4(3):221-31.
2. Biswas A, Oh PI, Faulkner GE et al. Sedentary time and its association with risk for disease incidence, mortality, and hospitalization in adults: a systematic review and meta-analysis. *Annals of Internal Medicine*. 2015;(2):123.
3. Thorp AA, Owen N, Neuhaus M, Dunstan DW. Sedentary Behaviors and Subsequent Health Outcomes in Adults: A Systematic Review of Longitudinal Studies, 1996–2011. *American Journal of Preventive Medicine*. 2011;41(2):207-15.
4. Yang Y, Lynch BM, van Roekel EH. Letter by Yang et al Regarding Article, "Accelerometer-Measured Physical Activity and Sedentary Behavior in Relation to All-Cause Mortality: The Women's Health Study". *Circulation*. 2018;138(1):114-5.
5. Pedisic Z, Grunseit A, Ding D et al. High sitting time or obesity: Which came first? Bidirectional association in a longitudinal study of 31,787 Australian adults. *Obesity*. 2014;22(10):2126-30.
6. Robins J. A new approach to causal inference in mortality studies with a sustained exposure period—application to control of the healthy worker survivor effect. *Mathematical Modelling*. 1986;7(9-12):1393-512.
7. Daniel R, Cousens S, De Stavola B, Kenward M, Sterne J. Methods for dealing with time-dependent confounding. *Statistics in Medicine*. 2013;32(9):1584-618.

8. Tager IB, Haight T, Sternfeld B, Yu Z, van Der Laan M. Effects of physical activity and body composition on functional limitation in the elderly: application of the marginal structural model. *Epidemiology*. 2004;15(4):479-93.
9. Bambom O, van der Laan M, Haight T, Tager I. Leisure-time physical activity and all-cause mortality in an elderly cohort. *Epidemiology*. 2009;20(3):424-30.
10. Haight T, Tager I, Sternfeld B, Satariano W, van der Laan M. Effects of Body Composition and Leisure-time Physical Activity on Transitions in Physical Functioning in the Elderly. *American Journal of Epidemiology*. 2005;162(7):607-17.
11. Garcia-Aymerich J, Lange P, Serra I, Schnohr P, Anto JM. Time-dependent confounding in the study of the effects of regular physical activity in chronic obstructive pulmonary disease: an application of the marginal structural model. *Annals of Epidemiology*. 2008;18(10):775-83.
12. Shortreed SM, Peeters A, Forbes AB. Estimating the effect of long-term physical activity on cardiovascular disease and mortality: evidence from the Framingham Heart Study. *Heart*. 2013;99(9):649-54.
13. Friedenreich CM, Neilson HK, Lynch BM. State of the epidemiological evidence on physical activity and cancer prevention. *European Journal of Cancer*. 2010;46(14):2593-604.
14. Westreich D. From Exposures to Population Interventions: Pregnancy and Response to HIV Therapy. *American Journal of Epidemiology*. 2014;179(7):797-806.
15. Dunstan DW, Zimmet PZ, Welborn TA et al. The Australian diabetes, obesity and lifestyle study (AusDiab)—methods and response rates. *Diabetes Research and Clinical Practice*. 2002;57(2):119-29.

16. *The Active Australia Survey: a guide and manual for implementation, analysis and reporting*. Australian Institute of Health and Welfare 2003. Available from: Australian Institute of Health and Welfare
17. Salmon J, Owen N, Crawford D, Bauman A, Sallis JF. Physical activity and sedentary behavior: A population-based study of barriers, enjoyment, and preference. *Health Psychology*. 2003;22(2):178-88.
18. Pink B. *An introduction to socio-economic indexes for areas (SEIFA)*. Canberra: Australian Bureau of Statistics. 2006.
19. Hodge A, English D, Itsiopoulos C, O'Dea K, Giles G. Does a Mediterranean diet reduce the mortality risk associated with diabetes: evidence from the Melbourne Collaborative Cohort Study. *Nutrition, Metabolism and Cardiovascular Diseases*. 2011;21(9):733-9.
20. Cerhan JR, Moore SC, Jacobs EJ et al. A pooled analysis of waist circumference and mortality in 650,000 adults. *Mayo Clinic Proceedings*. 2014;89(3):335-45.
21. Brown WJ, Bauman AE, Bull FC, Burton NW. Development of evidence-based physical activity recommendations for adults (18-64 years): report prepared for the Australian Government Department of Health, August 2012. 2013.
22. Wijndaele K, Healy GN, Dunstan DW et al. Increased cardiometabolic risk is associated with increased TV viewing time. *Medicine & Science in Sports & Exercise*. 2010;42(8):1511-8.
23. Taubman SL, Robins JM, Mittleman MA, Hernan MA. Intervening on risk factors for coronary heart disease: an application of the parametric g-formula. *International Journal of Epidemiology*. 2009;38(6):1599-611.

24. Lean ME, Han TS, Morrison CE. Waist circumference as a measure for indicating need for weight management. *British Medical Journal*. 1995;311(6998):158-61.
25. Jayasekara H, MacInnis RJ, Hodge AM et al. Alcohol consumption for different periods in life, intake pattern over time and all-cause mortality. *Journal of Public Health*. 2015;37(4):625-33.
26. White IR, Royston P, Wood AM. Multiple imputation using chained equations: Issues and guidance for practice. *Statistics in Medicine*. 2011;30(4):377-99.
27. Rubin DB. *Multiple imputation for nonresponse in surveys*. New York ; Wiley, c1987.; 1987.
28. Schomaker M, Heumann C. Bootstrap inference when using multiple imputation. *Statistics in Medicine*. 2018;37(14):2252-66.
29. Samitz G, Egger M, Zwahlen M. Domains of physical activity and all-cause mortality: systematic review and dose–response meta-analysis of cohort studies. *International Journal of Epidemiology*. 2011;40(5):1382-400.
30. Robins JM, Wasserman LA. Estimation of effects of sequential treatments by reparameterizing directed acyclic graphs. In: *Proceedings of the Thirteenth Conference on Uncertainty in Artificial Intelligence (UAI1997)*. 1997.
31. Hernán M, Robins J. *Causal Inference: What If*. Boca Raton: Chapman & Hall/CRC, forthcoming; 2020.
32. *Recommendations and public health and policy implications*. World Cancer Research Fund/American Institute for Cancer Research 2018. Available from: World Cancer Research Fund/American Institute for Cancer Research.

33. Global recommendations on physical activity for health. In: World Health Organization; 2010. <https://apps.who.int/iris/handle/10665/44399>
34. Australia's physical activity and sedentary behaviour guidelines. In: Australian Government Department of Health; 2014. <https://www1.health.gov.au/internet/main/publishing.nsf/Content/health-pubhlth-strateg-phys-act-guidelines>
35. Dunstan DW, Barr EL, Healy GN et al. Television viewing time and mortality: the Australian Diabetes, Obesity and Lifestyle Study (AusDiab). *Circulation*. 2010;121(3):384-91.
36. *National Health Survey 2004-05: Summary of Results*. Australian Bureau of Statistics 2006. Available from: Australian Bureau of Statistics
37. Chau J, Chey T, Burks-Young S, Engelen L, Bauman A. Trends in prevalence of leisure time physical activity and inactivity: results from Australian National Health Surveys 1989 to 2011. *Australian and New Zealand Journal of Public Health*. 2017;41(6):617-24.
38. VanderWeele TJ, Hernan MA. Causal Inference Under Multiple Versions of Treatment. *Journal of Causal Inference*. 2013;1(1):1-20.
39. VanderWeele TJ. On Well-defined Hypothetical Interventions in the Potential Outcomes Framework. *Epidemiology*. 2018;29(4):e24-e5.

Figure Caption

Figure 1. Plot of daily TV viewing and weekly physical activity at baseline (T1)

Figure 2. Effects of TV viewing and physical activity on all-cause mortality, estimated by g-formula and cox regression

The parametric g-formula (risk ratio and 95% CI) adjusted for time-fixed and time-varying confounding. Risk ratios for the '2 to <4 hours' TV viewing category and the '30-149 mins/week' physical activity category were not presented because we did not simulate these hypothetical interventions. Cox model (hazard ratio and 95% CI) adjusted for baseline (T1) confounders (i.e. typical adjustment of baseline confounders). Five imputations were used because of 5% T1 missing data. For comparison, '< 2 hours' TV viewing and '<30 mins/week' physical activity were used as reference categories in both the parametric g-formula analysis and the Cox model.

Supplemental Digital Content File

SDC 1 Figure.docx

SDC 2 Table.docx

SDC 3 Table.docx

SDC 4 Table.docx

SDC 5 Table.docx

Figure 1

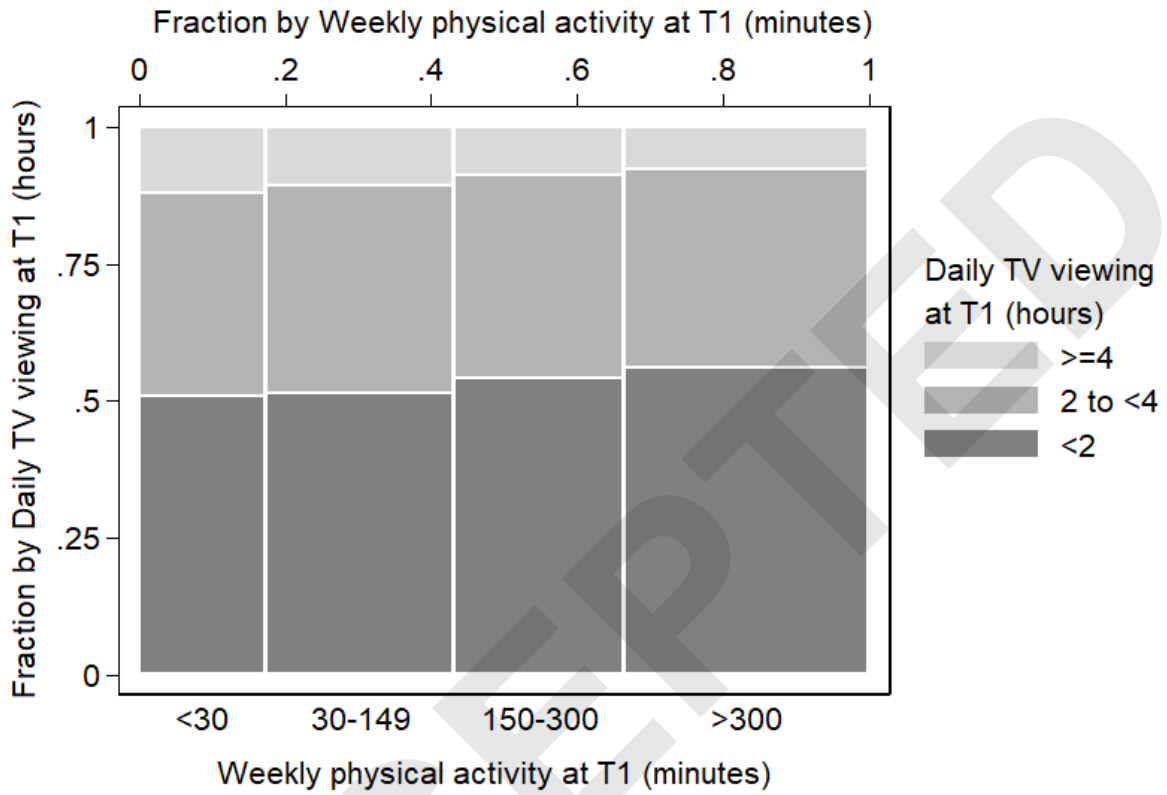


Figure 2

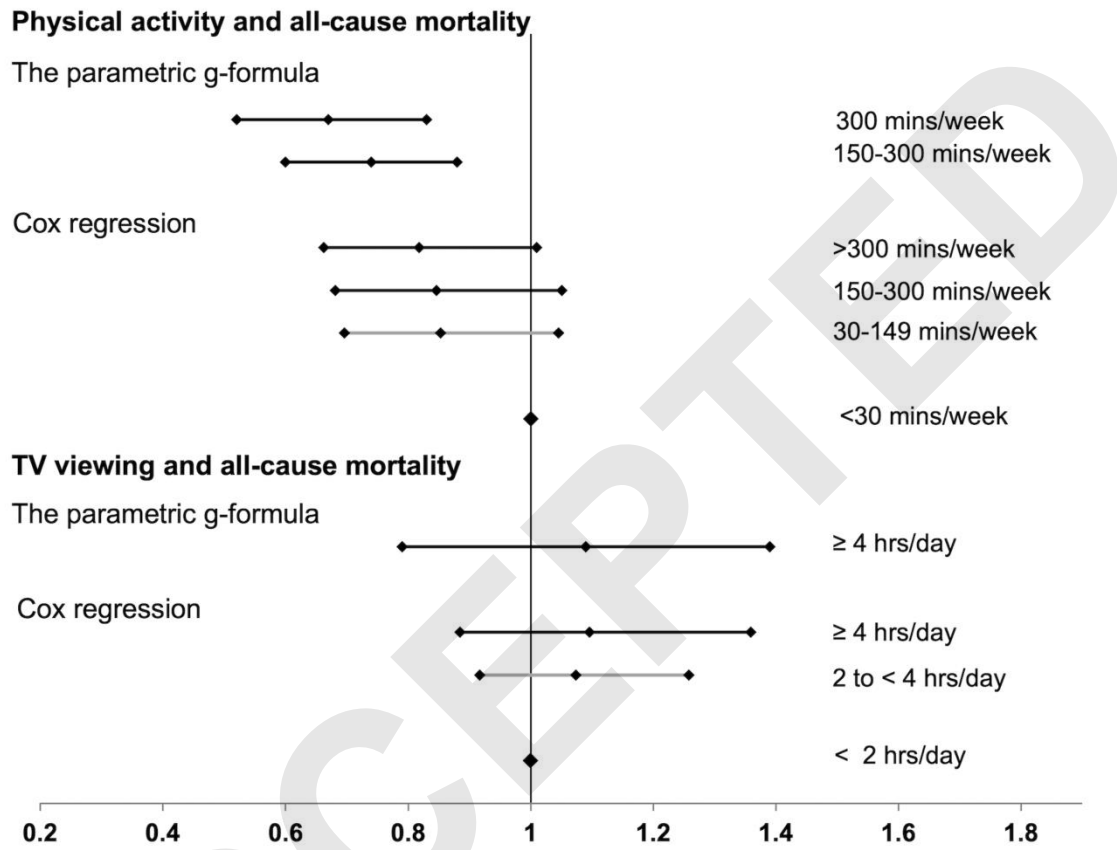


Table 1. Characteristics of participants included in the analysis, Australia

	1999-2000 (T0) N=6377	2004-05 (T1) N=6377	2011-12 (T2) N=4785 ^a
Time-fixed covariates			
Baseline age (years), mean(SD)		56.5 (12.8)	
Sex, N(%)			
Male		2891 (45.3)	
Female		3486 (54.7)	
Born in Australia/New Zealand, N(%)		4839 (75.9)	
The Index of Relative Socio-economic Advantage and Disadvantage (IRSAD), N(%)			
1 (greatest disadvantage)		1084 (17.3)	
2		1296 (20.7)	
3		1291 (20.6)	
4		1204 (19.2)	
5 (greatest advantage)		1395 (22.2)	
Level of education, N(%)			
University or technical institution		2561 (40.2)	
Completed high school		1460 (22.9)	
Some high school		1966 (30.8)	
Primary or never attended school		390 (6.1)	
Baseline height(cm), mean(SD)		167.6 (9.6)	
Time-varying covariates			
Weekly Physical activity, N(%)			
< 30 minutes	1257 (19.9)	1099 (17.4)	729 (15.8)
30 to 149 minutes	1686 (26.7)	1626 (25.7)	1127 (24.4)
150 to 300 minutes	1368 (21.6)	1480 (23.4)	1074 (23.3)
> 300 minutes	2015 (31.9)	2118 (33.5)	1680 (36.4)
Daily TV viewing time, N(%)			
< 2 hours	3655 (57.7)	3385 (53.6)	2030 (52.7)
2 to 4 hours	2225 (35.1)	2340 (37.0)	1478 (38.3)
≥ 4 hours	459 (7.2)	595 (9.4)	347 (9.0)
Mediterranean Diet Score, N(%)			
0-3	1870 (29.3)	1922 (30.7)	1067 (29.7)
4-6	3766 (59.1)	3695 (59.0)	2127 (59.3)
7-9	741 (11.6)	651 (10.4)	394 (11.0)
Waist circumference ^b , N(%)			
Normal	2500 (39.6)	2120 (33.3)	1057 (26.8)
Increased risk	1641 (26.0)	1654 (26.0)	1007 (25.5)
Greatly increased risk	2173 (34.4)	2584 (40.6)	1879 (47.7)

Table 1. Characteristics of participants included in the analysis, Australia (continued)

	1999-2000 (T0) N=6377	2004-05 (T1) N=6377	2011-12 (T2) N=4785 ^a
Smoking status, N(%)			
Never smoker	3686 (58.8)	3527 (58.0)	2657 (59.9)
Former smoker	1858 (29.6)	1982 (32.6)	1517 (34.2)
Current smoker	723 (11.5)	568 (9.3)	260 (5.9)
Alcohol intake (g/day), N(%)			
0 g/day (Male & Female)	940 (14.7)	836 (13.3)	481 (13.4)
1-39(Male)/1-19(Female)	4571 (71.7)	4470 (71.3)	2537 (70.7)
40-59(Male)/20-39(Female)	627 (9.8)	683 (10.9)	411 (11.5)
60+(Male)/40+(Female)	239 (3.7)	279 (4.5)	159 (4.4)
Self-reported general health, N(%)			
Excellent	603 (9.5)	689 (10.9)	426 (10.7)
Very Good	2346 (37.0)	2335 (36.9)	1522 (38.3)
Good	2633 (41.5)	2460 (38.8)	1552 (39.1)
Fair	693 (10.9)	755 (11.9)	422 (10.6)
Poor	74 (1.2)	95 (1.5)	51 (1.3)
History of health conditions, N(%)			
High cholesterol	1714 (27.0)	2654 (41.8)	3044 (58.1)
High blood pressure	1690 (26.6)	2399 (37.7)	2666 (51.5)
Diabetes	276 (4.3)	512 (8.0)	629 (12.9)
Heart conditions	443 (7.0)	559 (8.8)	218 (4.6)

Numbers across categories for some variables did not add up because of missing values.

^a Number of participants attended T2, before multiple imputation was applied to impute missing data due to missing T2 attendance for those who were still alive at T2. ^b Normal: <94cm (male) or <80cm (female); increased risk: 94cm to <102cm (male) or 80cm to <88 cm (female); greatly increased risk: ≥102cm (male) or ≥88 (women).

Table 2. Risks of death under hypothetical interventions using the parametric g-formula

No.	Interventions		13-year risk of death (%), 95% CI	Population risk difference (%), 95% CI	Population risk ratio, 95% CI	Risk difference (%), 95% CI	Risk ratio, 95% CI	Average % needed intervention ^a
0	Natural course	No intervention	12.1 (10.9 to 13.2)	Reference	Reference			0
1	Worst-case scenario	Reducing physical activity to <30 mins/week, and increasing TV viewing to ≥4 hrs/day for all	15.2 (11.6 to 18.9)	3.2 (-0.4 to 6.8)	1.26 (0.96 to 1.57)	Reference	Reference	97.6
2	Physical activity only, moderate	Increasing physical activity to 150-300 mins/week if <150 mins/week	11.1 (9.7 to 12.4)	-1.0 (-2.2 to 0.2)	0.92 (0.82 to 1.01)	-4.2 (-8.2 to -0.2)	0.73 (0.52 to 0.94)	42.1
3	Physical activity only, intensive	Increasing physical activity to >300 mins/week for all	10.0 (8.6 to 11.5)	-2.0 (-3.3 to -0.7)	0.83 (0.73 to 0.94)	-5.2 (-9.3 to -1.1)	0.66 (0.46 to 0.86)	65.2
4	TV viewing only	Reducing TV viewing to <2 hrs/day if ≥2 hrs/day	12.8 (11.1 to 14.6)	0.8 (-0.9 to 2.4)	1.06 (0.93 to 1.20)	-2.4 (-6.6 to 1.8)	0.85 (0.60 to 1.10)	48.4
5	Joint, moderate	Intervention No. 2 and No.4	11.6 (9.8 to 13.3)	-0.5 (-2.1 to 1.1)	0.96 (0.83 to 1.09)	-3.7 (-8.0 to 0.7)	0.76 (0.52 to 1.01)	68.2
6	Joint, intensive	Intervention No. 3 and No.4	10.5 (8.7 to 12.4)	-1.5 (-3.3 to 0.2)	0.87 (0.73 to 1.02)	-4.7 (-9.2 to -0.2)	0.70 (0.46 to 0.93)	80.7

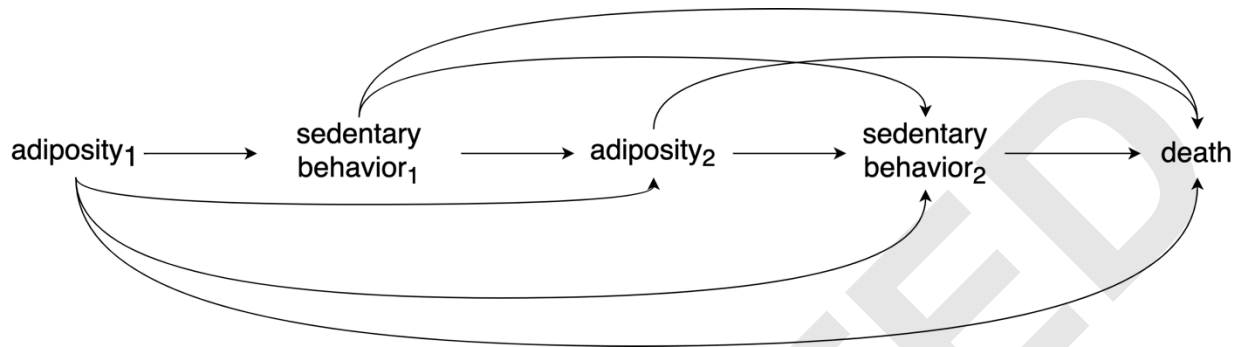
The observed 13-year risk of death was 12.2%; 500 bootstrap samples were drawn for each of the 40 imputed datasets to estimate the standard errors and 95% CIs. ^a Average percentage of participants who need to be intervened on at T1 and T2.

Table 3. Risk of death under hypothetical interventions in women and men

Interventions	13-year risk of death (%), 95% CI	Population risk ratio, 95% CI	Population risk difference (%), 95% CI	Risk ratio, 95% CI	Risk difference (%), 95% CI
Women					
Natural course	9.9 (8.4 to 11.3)	Reference	Reference		
Worst-case scenario	12.5 (8.0 to 16.9)	1.27 (0.83 to 1.70)	2.6 (-1.6 to 6.8)	Reference	Reference
Joint, intensive	8.7 (6.1 to 11.3)	0.88 (0.61 to 1.15)	-1.2 (-3.8 to 1.4)	0.70 (0.31 to 1.09)	-3.8 (-9.6 to 2.0)
Men					
Natural course	14.7 (13.1 to 16.4)	Reference	Reference		
Worst-case scenario	19.1 (12.9 to 25.2)	1.30 (0.88 to 1.71)	4.3 (-1.8 to 10.5)	Reference	Reference
Joint, intensive	12.7 (9.7 to 15.7)	0.86 (0.68 to 1.04)	-2.0 (-4.7 to 0.7)	0.67 (0.36 to 0.98)	-6.4 (-14.0 to 1.3)

The observed 13-year risk of death was 9.8% for women, and 15.2% for men; 200 bootstrap samples were drawn for each of the 40 imputed datasets to estimate the standard errors and 95% CIs.

SDC 1. Causal diagram - an example of time-varying confounding affected by prior exposure.



When estimating the effect of *sedentary behavior*₁ and *sedentary behavior*₂ on risk of *death* in a regression model, not conditioning on *adiposity*₂ will lead to unadjusted confounding. However, adjusting for *adiposity*₂ will block some causal pathways from *sedentary behavior*₁ to *death*.

SDC 2. Summary of variables used to model 13-year risk of death

Variables	Categories	Collection wave(s)	Type of model when used as dependent variable	Variables conditioned on in the model	Functional form when used as independent variable
Death between T1 and T2	Yes; No	Linkage	Logistic	All prior values of physical activity and TV viewing time, with an interaction term between the most recent values; past values of all time-varying covariates and comorbidity histories; and all time-fixed covariates.	-
Death between T2 and 14 April 2017	Yes; No	Linkage	Logistic	As above, for participants who are simulated to be alive at T2.	-
Weekly physical activity	< 30 minutes; 30 to 149 minutes; 150 to 300 minutes; > 300 minutes	T0, T1, T2	T0: not predicted T1 and T2: Ordered logistic	As above if not part of a hypothetical intervention; set to an intervening value if part of a hypothetical intervention.	Categories as pseudo-continuous
Daily TV viewing	< 2 hours; 2 to < 4 hours; ≥ 4 hours	T0, T1, T2	T0: not predicted T1 and T2: Ordered logistic	As above, if not part of a hypothetical intervention; set to an intervening value if part of a hypothetical intervention.	3 categories
Sex	Male; Female	T1	Time-fixed, not predicted	-	2 categories
Baseline age	Continuous in years	T1	Time-fixed, not predicted	-	Continuous
The Index of Relative Socio-economic Advantage and Disadvantage (IRSAD) quintiles	1 (greatest disadvantage); 2; 3; 4; 5 (greatest advantage)	T1	Time-fixed, not predicted	-	Quintile indicators
Born in Australia/New Zealand	Yes; No	T1	Time-fixed, not predicted	-	2 categories
Level of education	University or technical institution; completed high school; some high school; primary/never attended school	T1	Time-fixed, not predicted	-	4 categories
History of high cholesterol	Yes; No	T0, T1, T2	T0 and T1: not predicted T2: Logistic ^a	For participants simulated to be alive at T2: All prior values of physical activity and TV viewing time, with an interaction term between the most recent values; past values of all time-varying covariates and comorbidity histories; and all time-fixed covariates.	2 categories
History of high blood pressure	Yes; No	T0, T1, T2	T0 and T1: not predicted T2: Logistic ^a	As above.	2 categories
History of heart diseases	Yes; No	T0, T1, T2	T0 and T1: not predicted T2: Logistic ^a	As above.	2 categories

Supplementary Table 1. Summary of variables used to model 13-year risk of death (continued)

Variables	Categories	Collection wave(s)	Type of model when used as dependent variable	Variables conditioned on in the model	Functional form when used as independent variable
History of diabetes	Yes; No	T0, T1, T2	T0 and T1: not predicted T2: Logistic ^a	As above.	2 categories
Self-reported general health	Excellent; very good; good; fair; poor	T0, T1, T2	T0 and T1: not predicted T2: Ordered logistic	As above.	Categories as pseudo-continuous
Waist circumference ^b	Normal; increased risk; greatly increased risk	T0, T1, T2	T0 and T1: not predicted T2: Ordered logistic	As above.	3 categories
Mediterranean Diet Score	0 to 3; 3 to 6; 7 to 9	T0, T1, T2	T0 and T1: not predicted T2: Ordered logistic	As above.	3 categories
Smoking status	Never smoker; former smoker; current smoker	T0, T1, T2	T0 and T1: not predicted T2: Logistic	As above. Probability of quitting smoking at T2 was modelled for T1 current smokers. We assumed non-smokers would not start smoking at T2.	3 categories
Alcohol intake (g/day)	0 (Male & Female); 1-39 (Male)/1-19(Female); 40-59(Male)/20-39(Female); 60+(Male)/40+(Female) g/day	T0, T1, T2	T0 and T1: not predicted T2: Ordered logistic	As above.	Categories as pseudo-continuous

^a For these variables, T2 value was predicted based on a logistic regression only if self-reported 'No' at T1. ^b Normal: <94cm (male) or <80cm (female); increased risk: 94cm to <102cm (male) or 80cm to <88 cm (female); greatly increased risk: ≥102cm (male) or ≥88 (women).

SDC 3. T1 Characteristics of participants attended T2 and participants who were alive but did not attend T2

T1 Characteristics	Sample at T1		Attended T2		Did not attend T2	
	n=6377		n=4785		n=1219	
Baseline age (years), mean(SD)	56.5	(12.8)	54.5	(11.4)	59.4	(14.1)
Sex, N(%)						
Male	2891	(45.3)	2120	(44.3)	547	(44.9)
Female	3486	(54.7)	2665	(55.7)	672	(55.1)
Born in Australia/New Zealand, N(%)	4839	(75.9)	3725	(77.8)	843	(69.2)
The Index of Relative Socio-economic Advantage and Disadvantage (IRSAD), N(%)						
1 (greatest disadvantage)	1084	(17.3)	725	(15.4)	269	(22.4)
2	1296	(20.7)	937	(19.9)	273	(22.8)
3	1291	(20.6)	967	(20.5)	258	(21.5)
4	1204	(19.2)	940	(20.0)	203	(16.9)
5 (greatest advantage)	1395	(22.2)	1137	(24.2)	196	(16.3)
Level of education, N(%)						
University or technical institution	2561	(40.2)	2109	(44.1)	368	(30.2)
Completed high school	1460	(22.9)	1087	(22.7)	297	(24.4)
Some high school	1966	(30.8)	1413	(29.5)	422	(34.6)
Primary or never attended school	390	(6.1)	176	(3.7)	132	(10.8)
Baseline height(cm), mean(SD)	167.6	(9.6)	168.1	(9.4)	166.2	(9.9)
Weekly Physical activity, N(%)						
< 30 minutes	1099	(17.4)	763	(16.0)	227	(18.9)
30 to 149 minutes	1626	(25.7)	1239	(26.1)	299	(24.9)
150 to 300 minutes	1480	(23.4)	1112	(23.4)	278	(23.1)
> 300 minutes	2118	(33.5)	1642	(34.5)	398	(33.1)
Daily TV viewing time, N(%)						
< 2 hours	3385	(53.6)	2671	(56.1)	579	(48.3)
2 to 4 hours	2340	(37.0)	1707	(35.9)	467	(38.9)
>= 4 hours	595	(9.4)	379	(8.0)	154	(12.8)
Mediterranean Diet Score, N(%)						
0-3	1922	(30.7)	1438	(30.5)	369	(31.1)
4-6	3695	(59.0)	2770	(58.7)	706	(59.5)
7-9	651	(10.4)	513	(10.9)	112	(9.4)
Waist circumference*, N(%)						
Normal	2120	(33.3)	1665	(34.9)	365	(30.1)
Increased risk	1654	(26.0)	1238	(25.9)	312	(25.7)
Greatly increased risk	2584	(40.6)	1871	(39.2)	536	(44.2)

SDC 5. T1 Characteristics of participants attended T2 and participants who were alive but did not attend T2 (continued)

T1 Characteristics	Sample at T1		Attended T2		Did not attend T2	
	n=6377		n=4785		n=1219	
Smoking status, N(%)						
Never smoker	3527	(58.0)	2732	(59.6)	628	(54.7)
Former smoker	1982	(32.6)	1461	(31.9)	378	(32.9)
Current smoker	568	(9.3)	390	(8.5)	143	(12.4)
Alcohol intake (g/day), N(%)						
0 g/day (Male & Female)	836	(13.3)	527	(11.2)	220	(18.5)
1-39(Male)/1-19(Female)	4470	(71.3)	3418	(72.4)	806	(67.9)
40-59(Male)/20-39(Female)	683	(10.9)	549	(11.6)	115	(9.7)
60+(Male)/40+(Female)	279	(4.5)	227	(4.8)	46	(3.9)
Self-reported general health, N(%)						
Excellent	689	(10.9)	592	(12.4)	83	(6.9)
Very Good	2335	(36.9)	1878	(39.5)	371	(30.8)
Good	2460	(38.8)	1795	(37.7)	524	(43.5)
Fair	755	(11.9)	454	(9.5)	202	(16.8)
Poor	95	(1.5)	41	(0.9)	25	(2.1)
History of health conditions, N(%)						
High cholesterol	2654	(41.8)	1920	(40.2)	562	(46.5)
High blood pressure	2399	(37.7)	1663	(34.8)	507	(41.9)
Diabetes	512	(8.0)	323	(6.8)	129	(10.6)
Heart conditions	559	(8.8)	276	(5.8)	158	(13.0)

* Normal: <94cm (male) or <80cm (female); increased risk: 94cm to <102cm (male) or 80cm to <88 cm (female); greatly increased risk: ≥102cm (male) or ≥88 (women).

SDC 4. Estimated risk of death under various modelling orders of time-varying covariates

Interventions	13-year risk of death (%), 95% CI	Population risk ratio, 95% CI	Population risk difference (%), 95% CI
Order 1^a			
Natural course	12.3 (11.1 to 13.5)	Reference	Reference
Joint, intensive	10.5 (8.7 to 12.4)	0.86 (0.70 to 1.01)	-1.8 (-3.7 to 0.1)
Order 2^b			
Natural course	12.3 (11.3 to 13.3)	Reference	Reference
Joint, intensive	10.9 (9.0 to 12.7)	0.88 (0.74 to 1.02)	-1.5 (-3.2 to 0.3)
Order 3^c			
Natural course	12.1 (10.9 to 13.2)	Reference	Reference
Joint, intensive	10.5 (8.7 to 12.4)	0.87 (0.73 to 1.01)	-1.6 (-3.3 to 0.2)
Order 4^d			
Natural course	12.0 (10.9 to 13.1)	Reference	Reference
Joint, intensive	10.4 (8.5 to 12.4)	0.87 (0.72 to 1.02)	-1.6 (-3.3 to 0.2)

The observed 13-year risk of death was 12.2%; 200 bootstrap samples were drawn for each of the 40 imputed datasets to estimate the standard errors and 95% CIs.

^aOrder 1: TV viewing, physical activity, high cholesterol, high blood pressure, heart conditions, diabetes, self-reported general health, waist circumference, Mediterranean Diet Score, smoking status, and alcohol intake.

^bOrder 2: Physical activity, TV viewing, high cholesterol, high blood pressure, heart conditions, diabetes, Mediterranean Diet Score, smoking, alcohol intake, waist circumference, and self-reported general health.

^cOrder 3: Physical activity, TV viewing, high blood pressure, high cholesterol, diabetes, heart conditions, Mediterranean Diet Score, smoking, alcohol intake, waist circumference, and self-reported general health.

^dOrder 4: Physical activity, TV viewing, self-reported general health, waist circumference, Mediterranean diet score, smoking, alcohol intake, high cholesterol, high blood pressure, heart conditions, and diabetes.

SDC 5. Pre-baseline (T0) characteristics of participants, by T1 attendance status

T0 Characteristics	Sample at T0		Attended T1		Did not attend T1	
	n=11247		n=6377		n=4847	
Baseline age (years), mean(SD)	51.5	(14.5)	51.5	(12.8)	51.6	(16.4)
Sex, N(%)						
Male	5048	(44.9)	2891	(45.3)	2155	(44.5)
Female	6199	(55.1)	3486	(54.7)	2692	(55.5)
Born in Australia/New Zealand, N(%)	8434	(75.0)	4839	(75.9)	3579	(73.8)
The Index of Relative Socio-economic Advantage and Disadvantage (IRSAD), N(%)						
1 (greatest disadvantage)	2241	(20.2)	1084	(17.3)	1155	(24.1)
2	2342	(21.1)	1296	(20.7)	1041	(21.7)
3	2236	(20.2)	1291	(20.6)	939	(19.6)
4	2065	(18.6)	1204	(19.2)	857	(17.9)
5 (greatest advantage)	2203	(19.9)	1395	(22.2)	803	(16.7)
Level of education, N(%)						
University or technical institution	4097	(36.4)	2561	(40.2)	1522	(31.4)
Completed high school	2403	(21.4)	1460	(22.9)	937	(19.3)
Some high school	3935	(35.0)	1966	(30.8)	1966	(40.6)
Primary or never attended school	809	(7.2)	390	(6.1)	419	(8.6)
Baseline height(cm), mean(SD)	168.1	(9.6)	168.5	(9.5)	167.6	(9.8)
Weekly Physical activity, N(%)						
< 30 minutes	2397	(21.5)	1257	(19.9)	1135	(23.7)
30 to 149 minutes	2946	(26.4)	1686	(26.7)	1254	(26.2)
150 to 300 minutes	2395	(21.5)	1368	(21.6)	1025	(21.4)
> 300 minutes	3401	(30.5)	2015	(31.9)	1376	(28.7)
Daily TV viewing time, N(%)						
< 2 hours	6174	(55.3)	3655	(57.7)	2510	(52.4)
2 to 4 hours	4038	(36.2)	2225	(35.1)	1801	(37.6)
>= 4 hours	943	(8.5)	459	(7.2)	482	(10.1)
Mediterranean Diet Score, N(%)						
0-3	3538	(31.5)	1870	(29.3)	1657	(34.2)
4-6	6529	(58.1)	3766	(59.1)	2754	(56.8)
7-9	1180	(10.5)	741	(11.6)	436	(9.0)
Waist circumference*, N(%)						
Normal	4287	(38.7)	2500	(39.6)	1772	(37.5)
Increased risk	2832	(25.6)	1641	(26.0)	1187	(25.1)
Greatly increased risk	3949	(35.7)	2173	(34.4)	1772	(37.5)
Smoking status, N(%)						
Never smoker	6072	(55.0)	3686	(58.8)	2371	(50.0)
Former smoker	3218	(29.2)	1858	(29.6)	1357	(28.6)
Current smoker	1745	(15.8)	723	(11.5)	1017	(21.4)

SDC 4. Pre-baseline (T0) characteristics of participants, by T1 attendance status (continued)

T0 Characteristics	Sample at T0	Attended T1	Did not attend T1
	n=11247	n=6377	n=4847
Alcohol intake (g/day), N(%)			
0 g/day (Male & Female)	1861 (16.5)	940 (14.7)	919 (19.0)
1-39(Male)/1-19(Female)	7894 (70.2)	4571 (71.7)	3303 (68.1)
40-59(Male)/20-39(Female)	1048 (9.3)	627 (9.8)	420 (8.7)
60+(Male)/40+(Female)	444 (3.9)	239 (3.7)	205 (4.2)
Self-reported general health, N(%)			
Excellent	936 (8.4)	603 (9.5)	331 (6.9)
Very Good	3786 (33.9)	2346 (37.0)	1436 (29.8)
Good	4739 (42.4)	2633 (41.5)	2092 (43.5)
Fair	1538 (13.8)	693 (10.9)	843 (17.5)
Poor	184 (1.6)	74 (1.2)	109 (2.3)
History of health conditions, N(%)			
High cholesterol	2832 (25.4)	1714 (27.0)	1116 (23.4)
High blood pressure	3046 (27.2)	1690 (26.6)	1354 (28.1)
Diabetes	567 (5.0)	276 (4.3)	291 (6.0)
Heart conditions	938 (8.4)	443 (7.0)	495 (10.3)

* Normal: <94cm (male) or <80cm (female); increased risk: 94cm to <102cm (male) or 80cm to <88 cm (female); greatly increased risk: ≥102cm (male) or ≥88 (women).